

# Evaluation of bootstrap methods in nonlinear mixed-effects models

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PAGE MEETING

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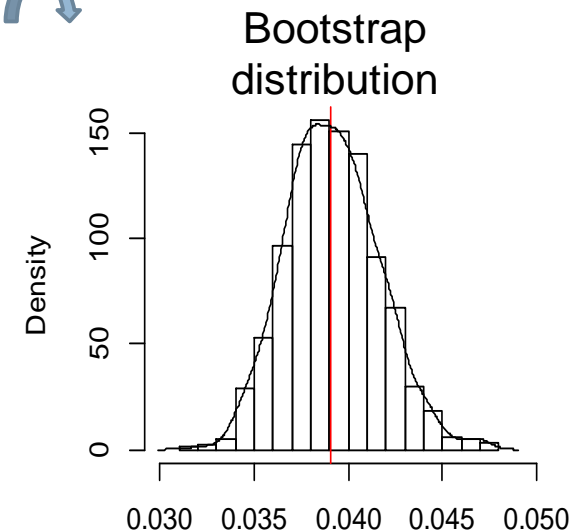
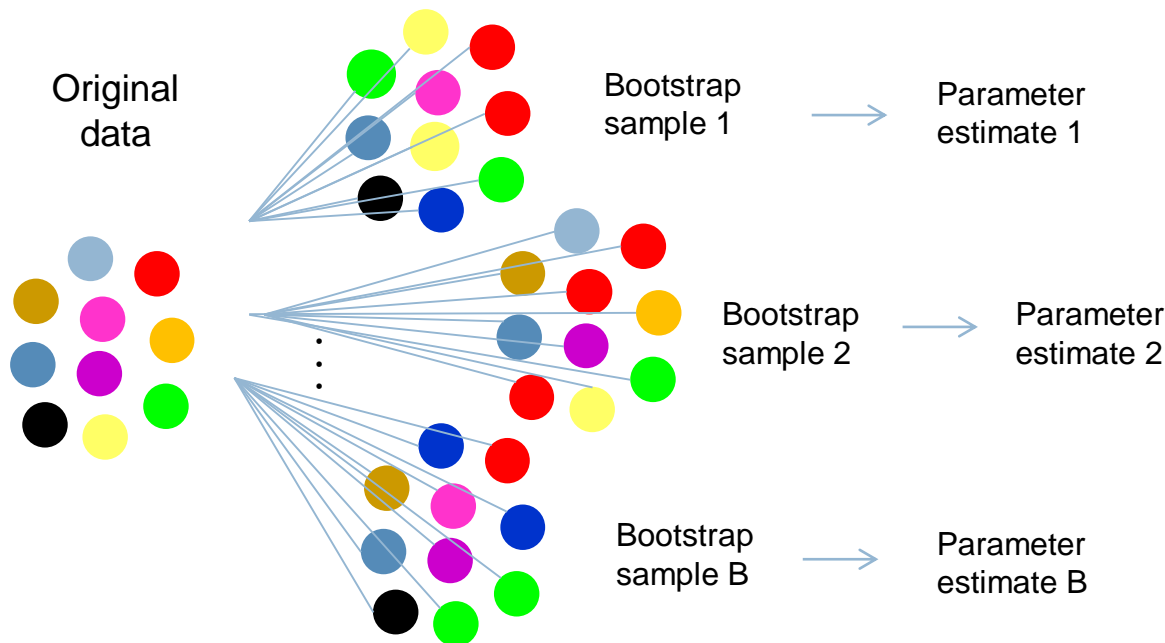




# Principle of bootstrap (Efron 1979)<sup>1</sup>

- Resample with replacement the observed data to construct the distribution of an estimate or a statistic of interest
- A general bootstrap procedure in regression:

«Pull one self up with your own bootstrap»



- “Main” bootstrap methods
  - ▣ case bootstrap<sup>2</sup>: resample the pairs
  - ▣ residual bootstrap<sup>1</sup>: resample the residuals after model fitting

1. Efron B. *Annal Stat* 1979; 7: 1-26

2. Freedman D.A. *Annal Stat* 1981; 9: 1218-1228

# Bootstrap in mixed-effects models (MEM)

## MEM model

$$y_i = f(\xi_i, \mu, \eta_i) + g(\xi_i, \mu, \eta_i, \sigma)\varepsilon_i$$

$\mu$ : fixed parameters

$$\eta_i \sim N(0, \Omega)$$

$$\varepsilon_i \sim N(0, 1)$$

- Take into account the characteristics of MEM<sup>1</sup>: repeated measures, heteroscedasticity, nonlinearity
- Respect two levels of variability: interindividual and residual variabilities<sup>2,3</sup>

## □ Previous simulation study in linear MEM to compare bootstrap methods

([THAI HT et al. Pharm Stat 2013; 12 \(3\): 129-140](#))

- ▣ poor performance of bootstraps resampling only one level of variability
- ▣ good performance of three bootstrap methods resampling two levels of variability
- ▣ some differences between bootstrap methods when applied to a real unbalanced dataset

1. Das S, Krishen A. *J Stat Plan Inference* 1999; 75: 237-245

2. Ocana J, Halimi E. *Mathematics Preprint Series* 2005; 367

3. Leeuw J, Meijer E. *Handbook of multilevel analysis*. Springer 2008

# Objective

Evaluate the performance of bootstrap methods for estimating uncertainty of parameters in nonlinear mixed-effects models (NLMEM) using a simulation study

# Bootstrap methods

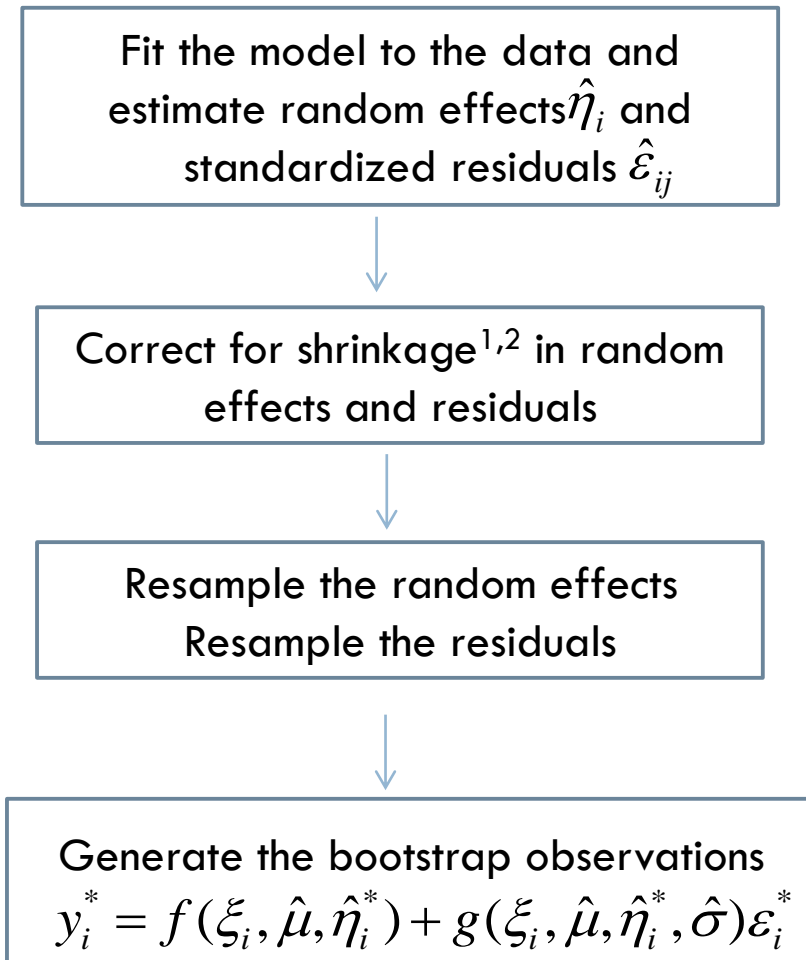
- **Case bootstrap ( $B_{\text{case}}$ )**
- Nonparametric random effect and residual bootstrap ( $B_{\eta,\varepsilon}^{\text{NP}}$ )
- Parametric random effect and residual bootstrap ( $B_{\eta,\varepsilon}^{\text{P}}$ )

Resample the individuals

$$(\xi_i, y_i) \rightarrow (\xi_i^*, y_i^*)$$

# Bootstrap methods

- Case bootstrap ( $B_{\text{case}}$ )
- **Nonparametric random effect and residual bootstrap ( $B_{\eta, \varepsilon}^{\text{NP}}$ )**
- Parametric random effect and residual bootstrap ( $B_{\eta, \varepsilon}^{\text{P}}$ )



1. Carpenter JR. *Appl Statist* 2003, 52: 431-443

2. Wang J et al. *Comput Methods Programs Biomed* 2006; 82: 130-143.

# Bootstrap methods

- Case bootstrap ( $B_{\text{case}}$ )
- Nonparametric random effect and residual bootstrap ( $B_{\eta,\varepsilon}^{\text{NP}}$ )
- **Parametric random effect and residual bootstrap ( $B_{\eta,\varepsilon}^{\text{P}}$ )**

Fit the model to the data



Simulate the random effects from  $N(0, \hat{\Omega})$   
Simulate the residuals from  $N(0, 1)$



Generate the bootstrap observations  
$$y_i^* = f(\xi_i, \hat{\mu}, \hat{\eta}_i^*) + g(\xi_i, \hat{\mu}, \hat{\eta}_i^*, \hat{\sigma}) \varepsilon_i^*$$

# Motivating example

## □ Pharmacokinetic data: aflibercept<sup>1</sup>, an anti-angiogenic agent

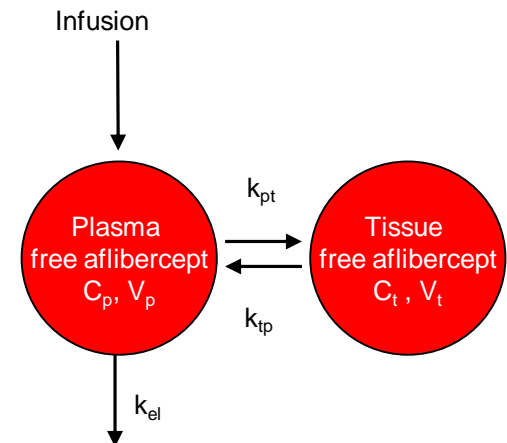
### □ 2 clinical trials<sup>2</sup>

- phase I TCD 6120 trial (first dose) ( $N_1=53$ ,  $n_1$ =median of 9)
- phase III VITAL trial (first two doses) ( $N_2=291$ ,  $n_2=2$ )

### □ two - compartment PK model with first-order elimination

- exponential model for IIV
- proportional model for residual error

## □ Model fit to the data: SAEM (MONOLIX 4.1.2)<sup>1</sup>



| Parameter             | Estimate (RSE) | IIV (%)   |
|-----------------------|----------------|-----------|
| CL (L/hr)             | 0.04 (2)       | 28.8 (5)  |
| V1 (L)                | 3.62 (2)       | 19.7 (10) |
| Q (L/hr)              | 0.14 (15)      | 111(12)   |
| V2 (L)                | 2.9 (5)        | -         |
| Corr(CL,Q) ( $\rho$ ) | 0.90 (8)       |           |
| $\sigma_p$ (%)        | 24.8 (4)       |           |

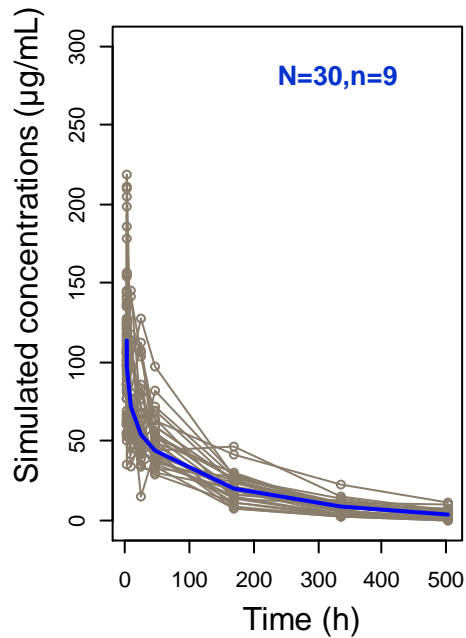
1. Chu QS. *Expert Opin Biol Ther* 2009;9:263-271  
 2. Gaya A et al. *Cancer Treat Rev* 2012 38): 484-493

3. Lavielle M. MONOLIX. MONOLIX group, Orsay, France, 2008.

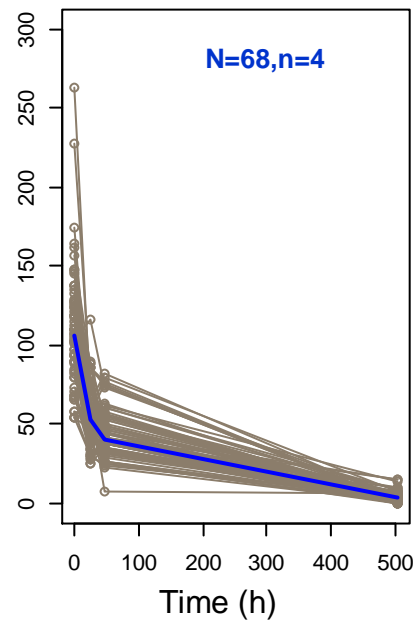


# Simulation settings

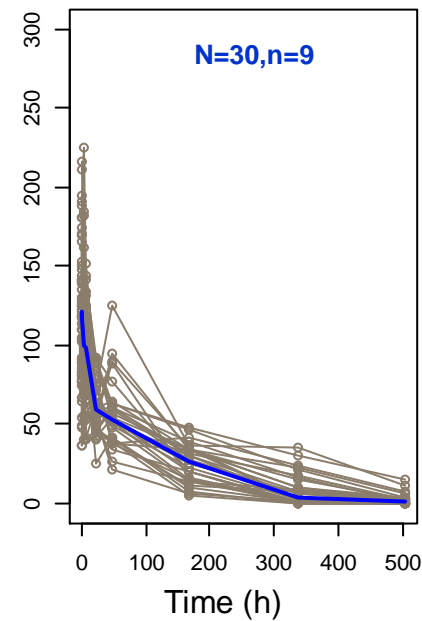
**Frequent balanced design**  
first-order



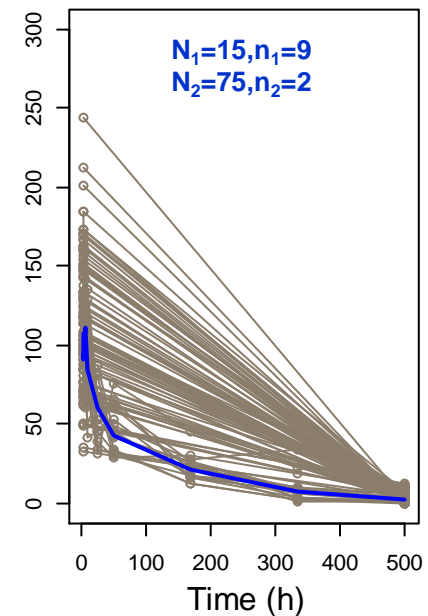
**Sparse balanced design**  
first-order



**Frequent balanced design**  
mixed-order (Michaelis-Menten)

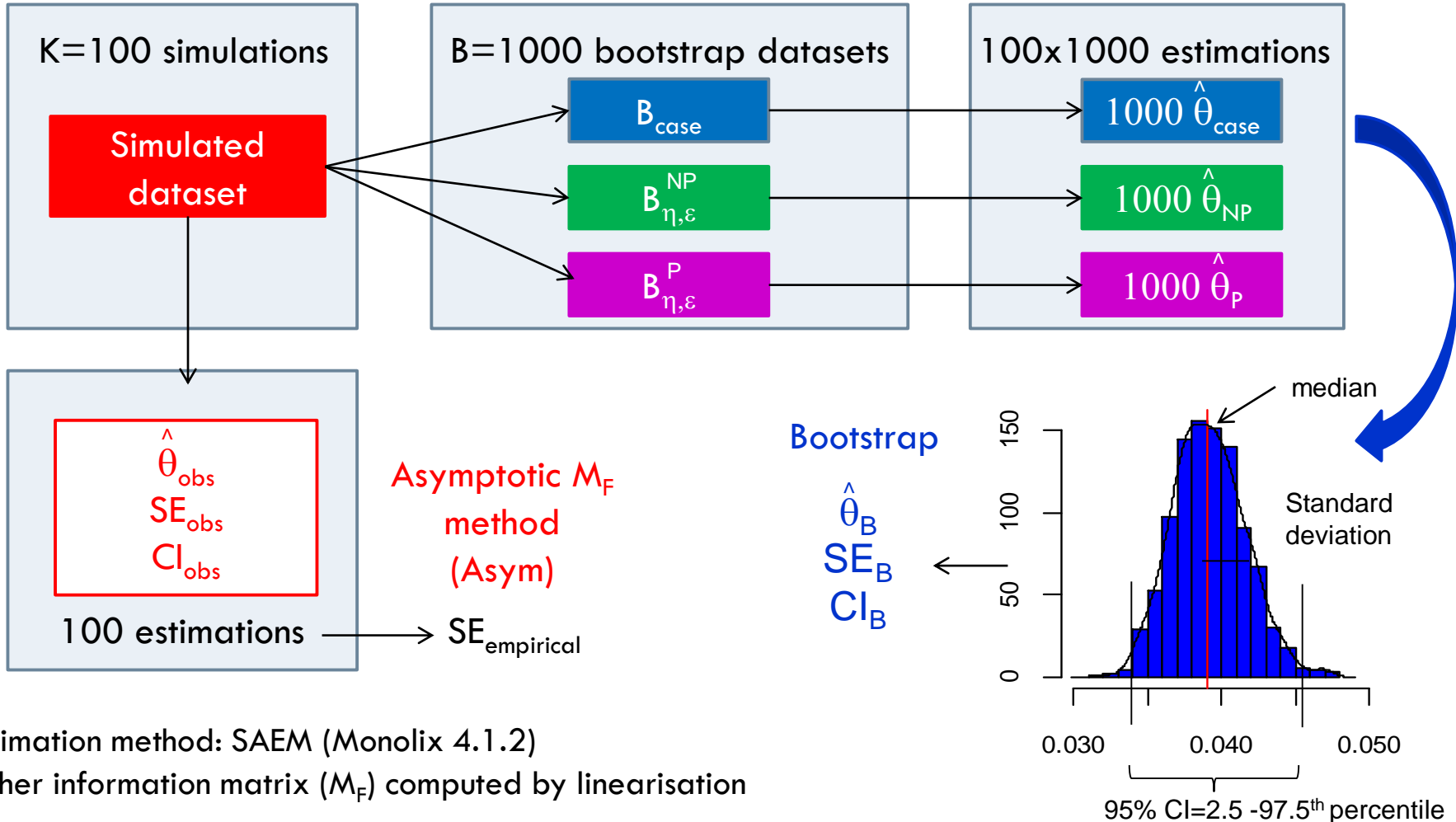


**Unbalanced design**  
first-order

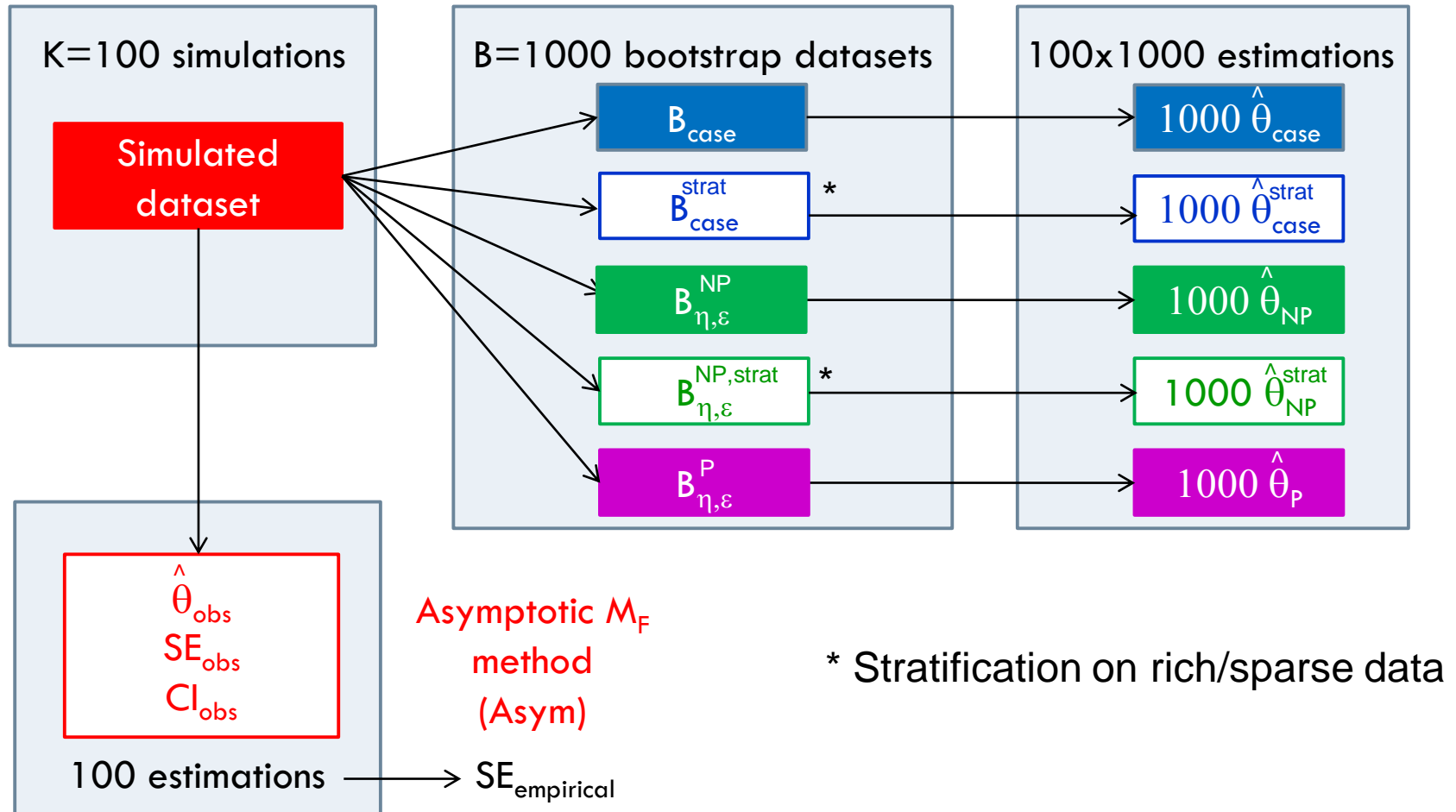


**K=100 replications for each design**

# Bootstrap settings



# Bootstrap settings for the unbalanced design



# Evaluation

- For each method (k=100 replicated datasets)

- relative bias of parameter (%)

$$\text{RBias}(\hat{\theta}_B^{(l)}) = \frac{1}{K} \sum_{k=1}^K \left( \frac{\hat{\theta}_{B;k}^{(l)} - \hat{\theta}_k^{(l)}}{\hat{\theta}_k^{(l)}} \times 100 \right)$$

- relative bias of SE (%)

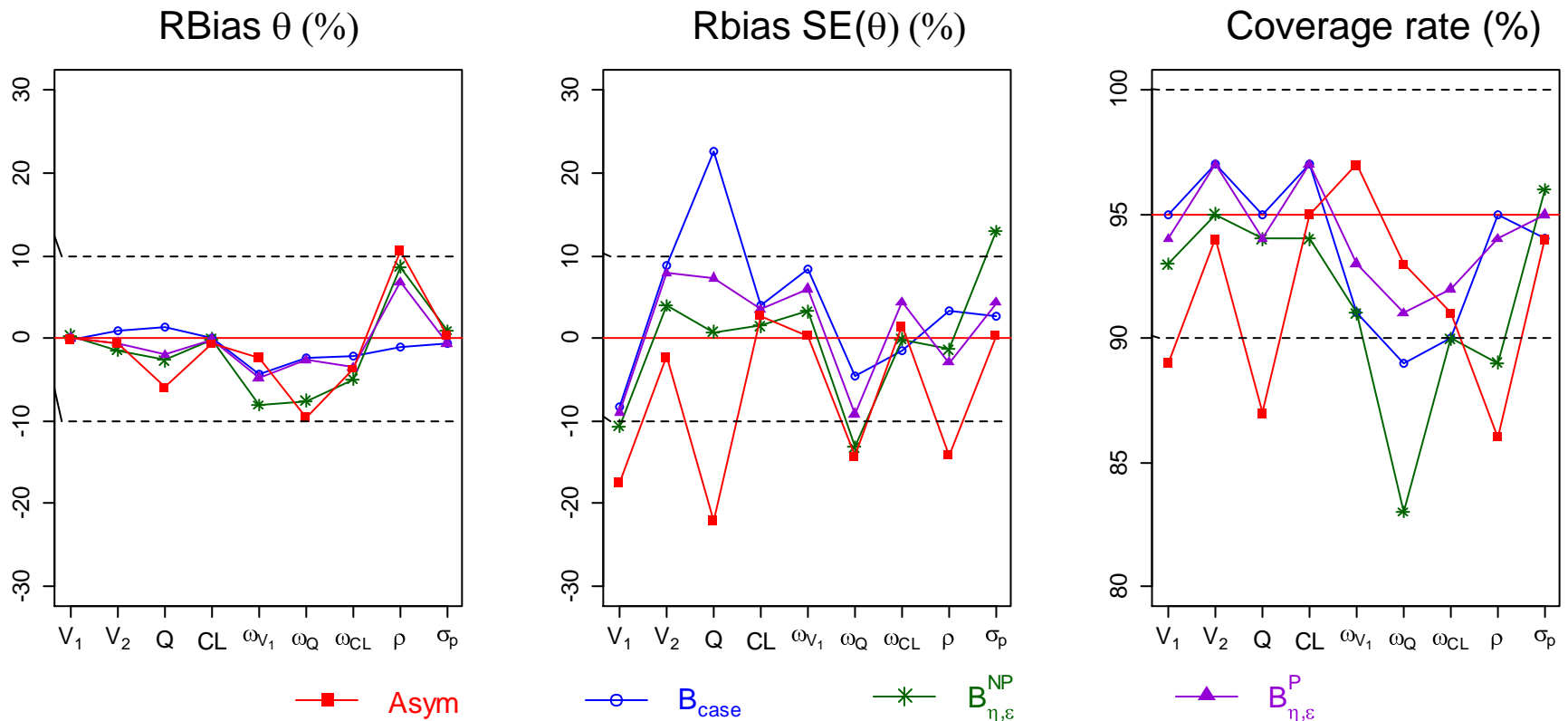
$$\text{RBias}(\widehat{SE}_B^{(l)}) = \frac{\frac{1}{K} \sum_{k=1}^K \widehat{SE}_{B;k}^{(l)} - \widehat{SE}_{\text{empirical}}^{(l)}}{\widehat{SE}_{\text{empirical}}^{(l)}} \times 100$$

- coverage rate of 95% CI: % bootstrap CI contain the true value of parameter

- **“Good”** bootstrap

- relative bias of parameters and SE: unbiased ( $\pm 10\%$ )
- coverage rate of 95% CI: good (90-100%)

# Frequent balanced design with first-order elimination



- ✓ **Parametric bootstrap** had the best performance
- ✓ **Asymptotic** approach performed slightly less well than the bootstrap methods

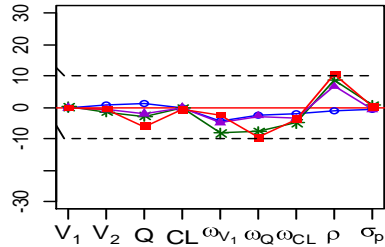
# Three balanced designs

Frequent design

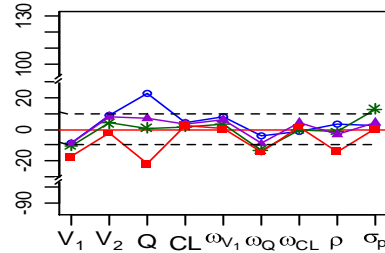
first order

(N=30, n=9)

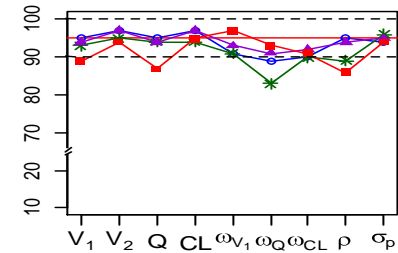
RBias  $\theta$  (%)



Rbias SE( $\theta$ ) (%)



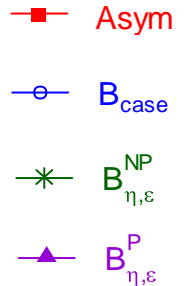
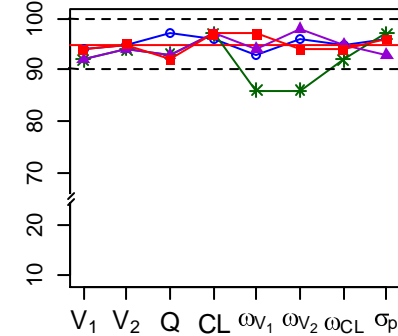
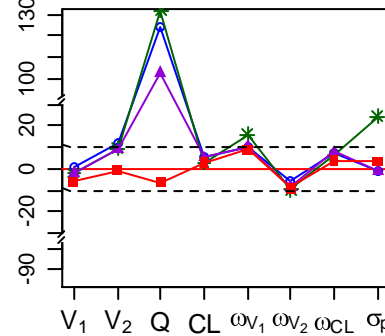
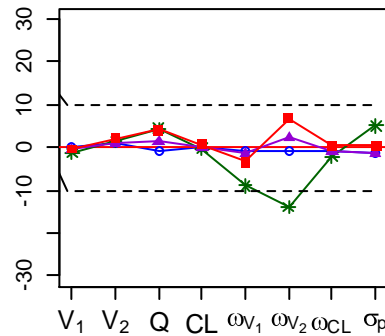
Coverage rate (%)



Sparse design

first order

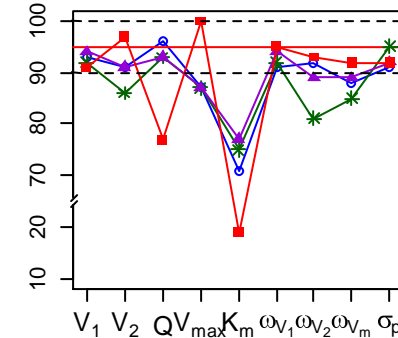
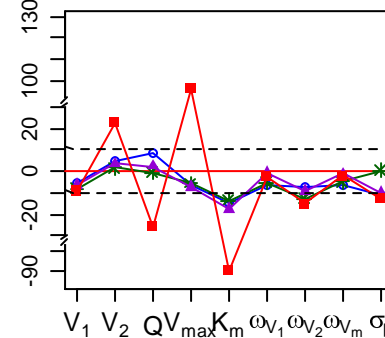
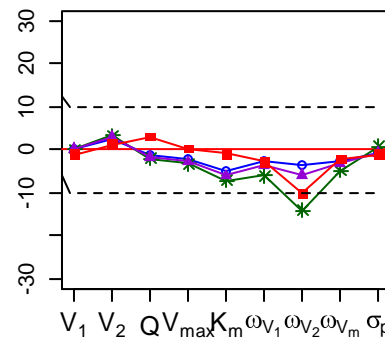
(N=68, n=4)



Frequent design

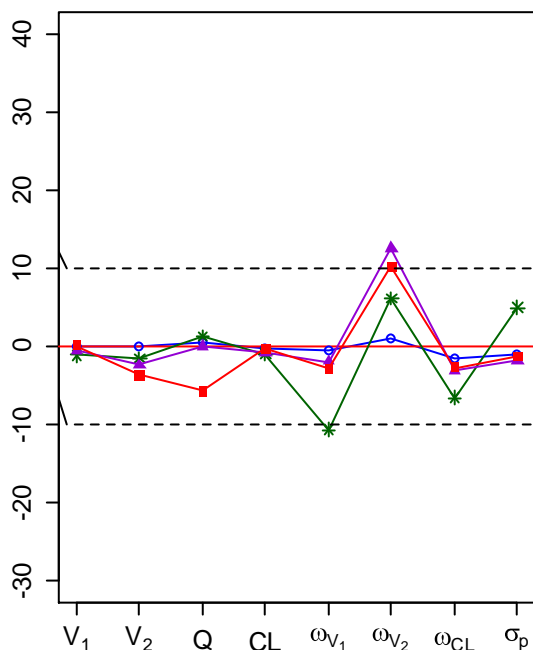
mixed-order

(N=30, n=9)



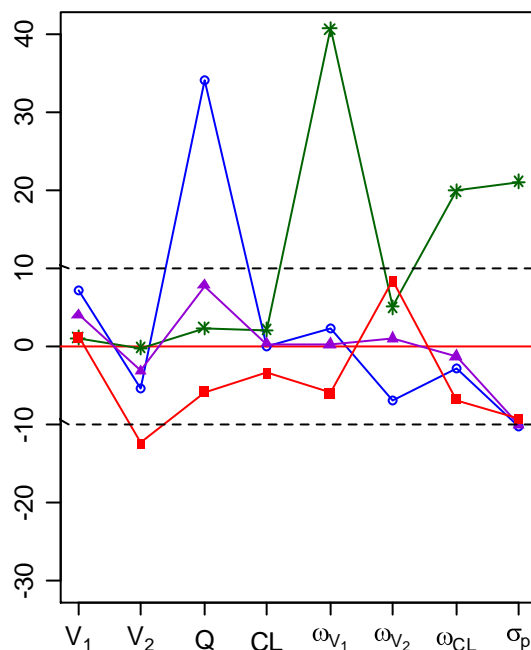
# Unbalanced design with first-order elimination

RBias  $\theta$  (%)



■ Asym

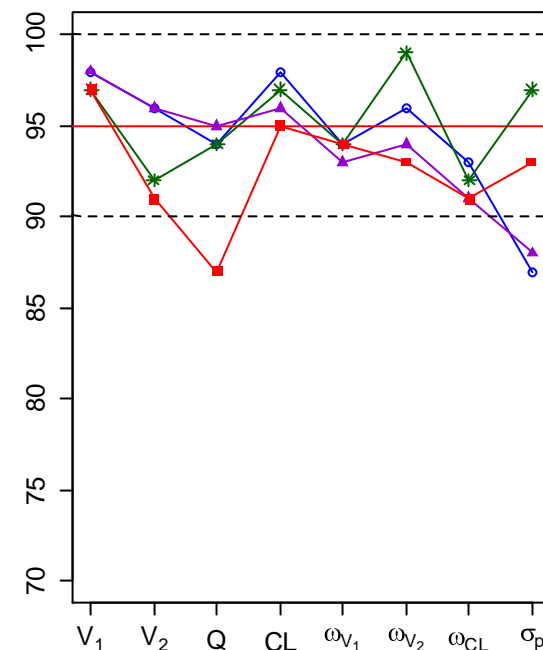
Rbias SE( $\theta$ ) (%)



○  $B_{case}$

\*  $B_{\eta,\varepsilon}^{NP}$

Coverage rate (%)

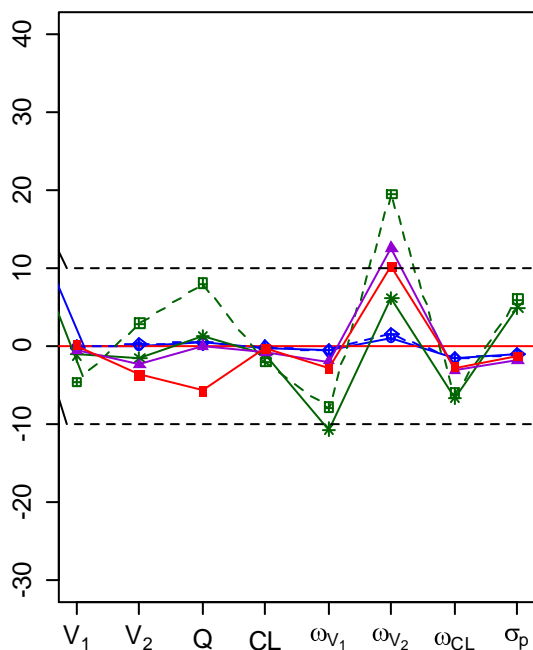


▲  $B_{\eta,\varepsilon}^P$

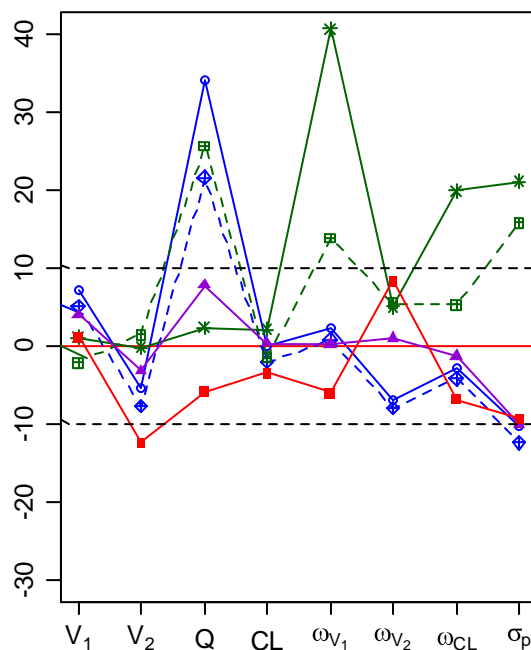
✓ **Asymptotic** approach and **parametric bootstrap** performed reasonably well and better than other bootstrap methods

# Unbalanced design with first-order elimination

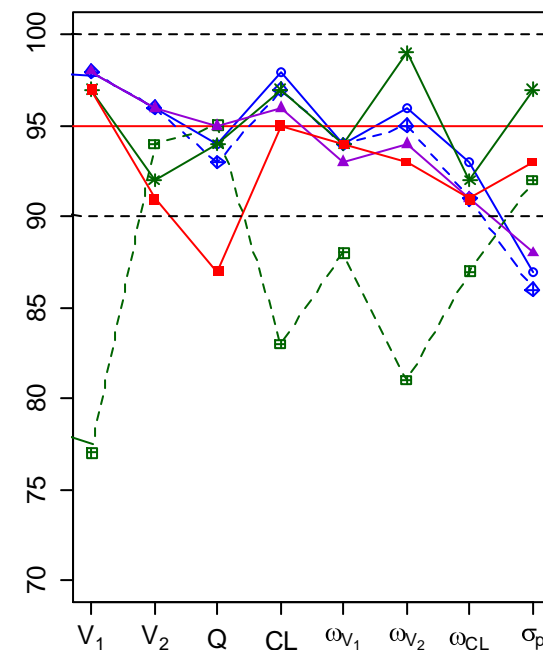
RBias  $\theta$  (%)



Rbias SE( $\theta$ ) (%)



Coverage rate (%)



—■— Asym

—○—  $B_{case}$

—\*—  $B_{\eta,\epsilon}^{NP}$

—▲—  $B_{\eta,\epsilon}^P$

—◇—  $B_{case}^{strat}$

—■—  $B_{\eta,\epsilon}^{NP, strat}$

✓ Stratification improved slightly the performance of **case bootstrap** and but degraded that of **nonparametric random effect and residual bootstrap** in term of coverage rate



# Conclusions

- Better estimation of uncertainty by the bootstrap methods than the **asymptotic** method in NLMEM with high nonlinearity
- Caution with bootstrap methods in presence of large fluctuations in parameter estimates between bootstrap samples
- The choice of bootstrap methods in NLMEM
  - ▣ **parametric bootstrap**: best description of uncertainty
    - study robustness in case of model misspecification
  - ▣ **case bootstrap**: fast, simple and robust (e.g heteroscedasticity, missing data)
    - evaluate stratification for complex designs
  - ▣ **nonparametric random effect and residual bootstrap**: maintain the same design as the original dataset
    - improve correction for shrinkage in unbalanced designs

# Acknowledgements

- Sanofi for financial support





# Annexes

# Correction for variance underestimation

## □ Correction for random effects<sup>1,2</sup>

□ center:  $\tilde{\eta}_i = \hat{\eta}_i - \bar{\eta}_i$

□ transform: by the ratio between empirical vs estimated variance-covariance ( $A_\eta$ )

■ Cholesky decomposition: matrix  $\Omega_R$  positive

■ EVD (*Eigen Value Decomposition*): matrix  $\Omega_R$  semi-positive

$$\hat{\eta}'_i = \tilde{\eta}_i \times A_\eta$$

## □ Correction for residuals<sup>1,2</sup>

□ center:  $\tilde{\epsilon}_{ij} = \hat{\epsilon}_{ij} - \bar{\epsilon}_{ij}$

□ transform: by the ratio between empirical vs estimated variance-covariance ( $A_\sigma$ )

■ homoscedastic error:  $A_\sigma = \hat{\sigma} / sd(\tilde{\epsilon}_{ij})$

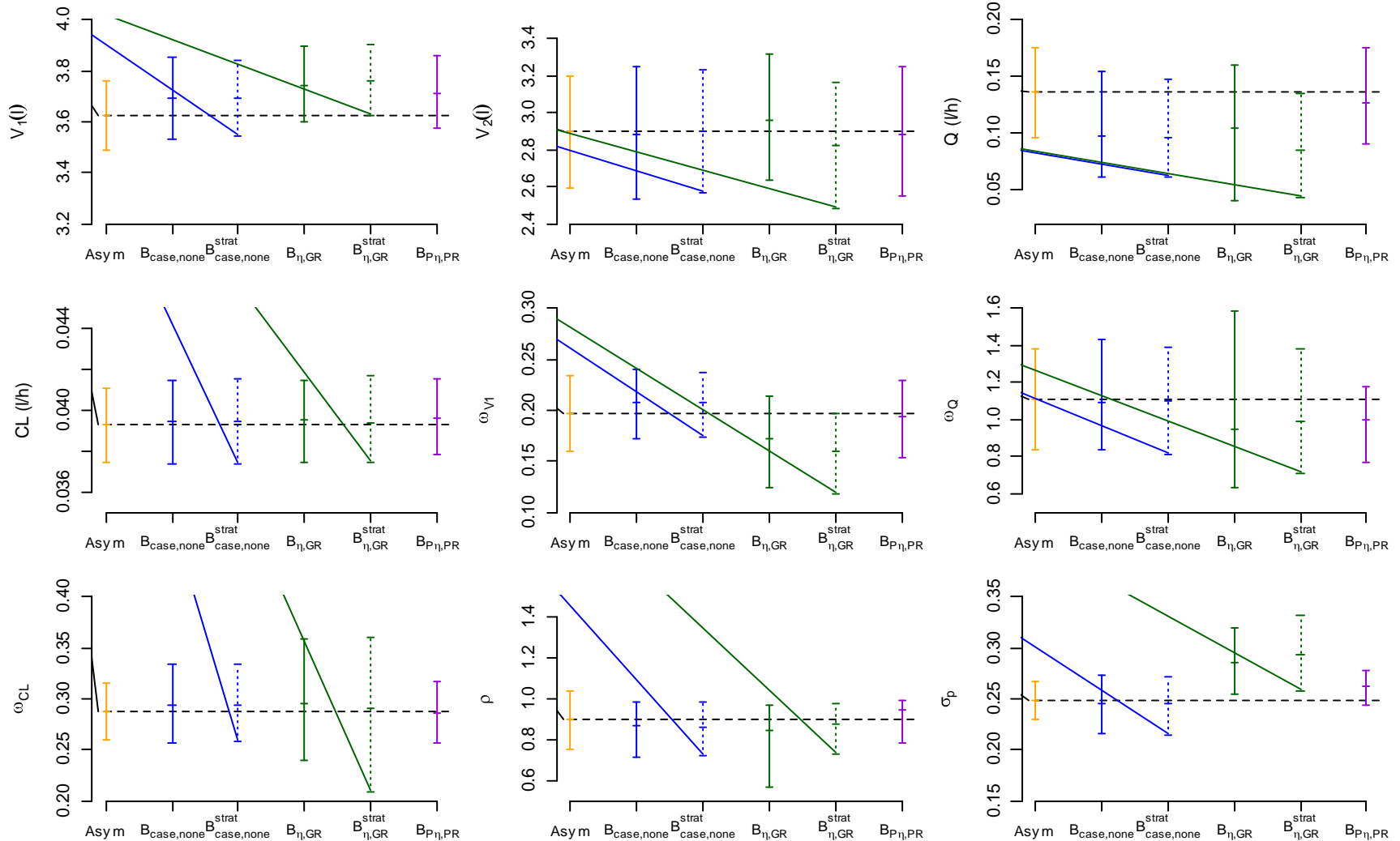
■ heteroscedastic error:  $A_\sigma = 1 / SD(\text{standardized residuals})$

$$\hat{\epsilon}'_{ij} = \tilde{\epsilon}_{ij} \times A_\sigma$$

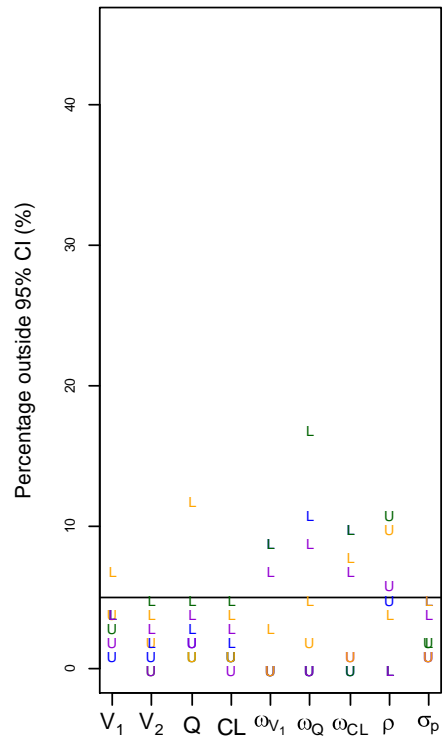
1. Carpenter JR. *Appl Statist* 2003, 52: 431-443

2. Wang J et al. *Comput Methods Programs Biomed* 2006; 82: 130-143.

# Application to aflibercept data

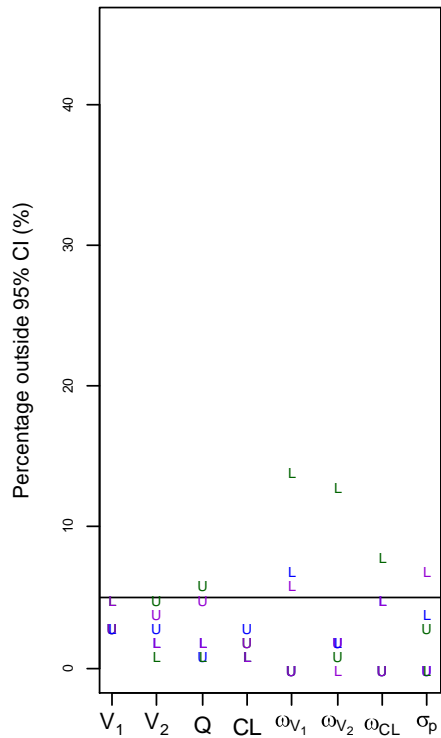


Frequent first-order balanced



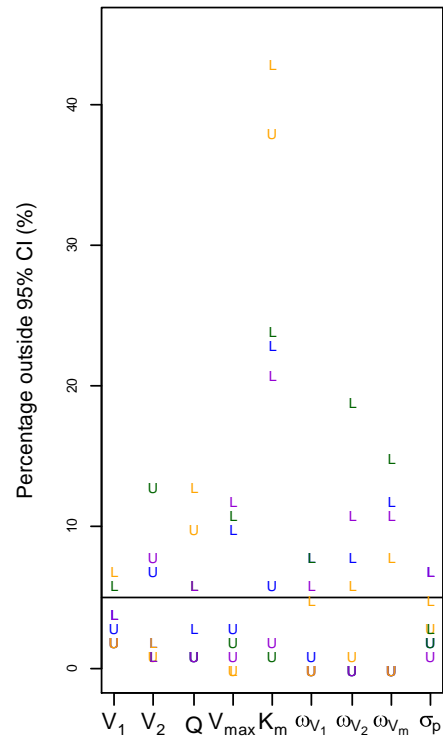
Asym

Sparse first-order balanced



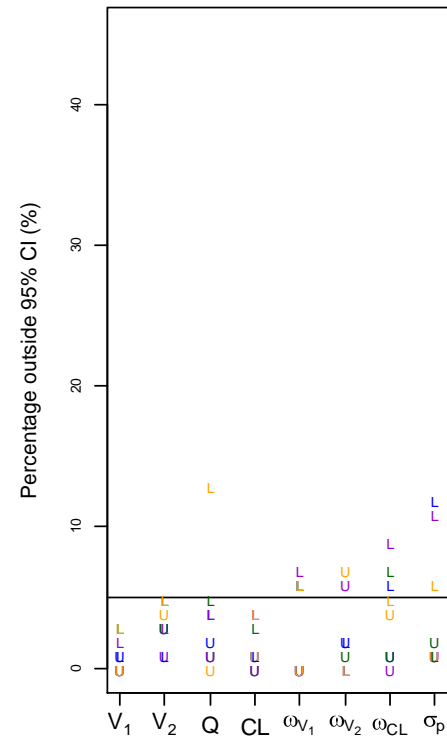
$B_{\text{case,none}}$

Frequent MM balanced



$B_{\eta,GR}$

First-order unbalanced



$B_{P_{\eta,PR}}$